

This is the "final" version. Revisions to the 1/21/00 draft are marked with asterisks: \*\*\*.

1. 2-32, p. 76, L-G

The number of distinct ordered triplets =  $60 \times 60 \times 60 = 60^3 = 216,000$

2. 2-35, p. 76, L-G

The number of distinct 7-tuples =  $8 \times 10 \times 10 \times 10 \times 10 \times 10 \times 10 = 8 \times 10^6$

3. 2-38, p. 76, L-G

There are  $3!$  permutations (possible orderings of 3 books) of which only one corresponds to the correct order. Assuming all permutations are equiprobable,

$$P(\text{correct order}) = \frac{1}{3!} = \frac{1}{6}$$

4. 2-40, p. 77, L-G

number of ways of picking 1 out of 5 =  $\binom{5}{1} = 5$

number of ways of picking 2 out of 5 =  $\binom{5}{2} = 10$

number of ways of picking none, some or all 5 =  $\sum_{i=1}^5 \binom{5}{i} = 2^5$  (by binomial thm)

5. 2-43, p. 77, L-G

There are  $\binom{49}{6}$  ways of choosing six numbers. Assuming each is equiprobable, the probability of winning is

$$\frac{1}{\binom{49}{6}} = 7.15 \times 10^{-8}$$

6. 2-54, p. 78, L-G (Note: Corollary 1 is in Section 2.2.)

The situation described is equivalent to the experiment of choosing twenty numbers in sequence from the set  $\{1, 2, \dots, 365\}$  with replacement and recording the 20 numbers in order.

A probability model for the experiment is:

Variable representing the outcome:  $\underline{X} = (X_1, \dots, X_{20})$ .

Sample space:  $S = \{\underline{x} = (x_1, \dots, x_{20}) : x_i \in \{1, \dots, 365\}\}$

Probability law:  $P(\underline{x}) = \frac{1}{365^{20}}$  for  $\underline{x} \in S$

(therefore, for any set  $A$ ,  $P(A) = |A| \times \frac{1}{365^{20}}$ )

We are interested in the probability of the event  $A$  that 2 or more students have the same birthday. As suggested by the hint, it is easier to find the probability of the complement  $A^c$ , which is the event that no two students have the same birthday.

$$A^c = \{ \underline{x} = (x_1, \dots, x_{20}) : x_i \in \{1, \dots, 365\}, x_i \neq x_j, i = 1, \dots, 20, j = 1, \dots, 20 \}$$

The number of elements in  $A^c$  is  $365 \times 364 \times \dots \times 346 = \frac{365!}{345!} = 1.037 \times 10^{51}$

$$\text{And so } P(A^c) = |A^c| \times \frac{1}{365^{20}} = \frac{365!}{345!} \times \frac{1}{365^{20}} = .589$$

$$\Rightarrow P(A) = 1 - P(A^c) = .411$$

If we repeat the problem with  $n = 50$  birthdays, then  $P(A) = .97$ .

7. A circular pan of water with radius 10 inches is stirred and the position of a floating spec is recorded after the water has come to rest.

(a) Determine the conditional probability that the spec is more than 1 inch from the center given that it is less than the 3 inches from the center.

We begin with a probability model:

Variable representing outcome:  $(X, Y)$  where  $X$  and  $Y$  represent the horizontal and vertical positions, respectively, of the spec after it comes to rest.

Sample space:  $S = \{ (x, y) : x^2 + y^2 \leq 10 \}$

$$\text{Probability law: } P(A) = \frac{\text{area } A}{\text{area } S} = \frac{\text{area } A}{100\pi}$$

(here we have given a general formula for the probability of any set)

We need to find  $P(A|B)$  where  $A = \{ (x, y) : 1 < x^2 + y^2 \leq 10 \}$  and  $B = \{ (x, y) : x^2 + y^2 < 3 \}$

From the definition of conditional probability

$$*** \quad P(A|B) = \frac{P(A \cap B)}{P(B)} = \frac{\text{area}(A \cap B)/100\pi}{\text{area}(B)/100\pi} = \frac{\text{area}(A \cap B)}{\text{area}(B)} = \frac{\pi 3^2 - \pi 1^2}{\pi 3^2} = \frac{8}{9}$$

(b) Determine the conditional probability that the spec is both to the north and east of the landing point, given that it is less than 5 inches from the center.

We need to find  $P(A|B)$  where  $A = \{ (x, y) : x^2 + y^2 \leq 10, x > 0, y > 0 \}$  and  $B = \{ (x, y) : x^2 + y^2 < 5 \}$

From the definition of conditional probability

$$P(A|B) = \frac{P(A \cap B)}{P(B)} = \frac{\text{area}(A \cap B)/100\pi}{\text{area}(B)/100\pi} = \frac{\text{area}(A \cap B)}{\text{area}(B)} = \frac{\pi 5^2/4}{\pi 5^2} = \frac{1}{4}$$

8. 2-49, p. 77, L-G.

Let  $B$  be a set such that  $P(B) > 0$ , so the axioms of probability are satisfied. Then,

Axiom 1:  $P(A|B) \geq 0$  for every set  $A$ , because  $P(A|B) = \frac{P(A \cap B)}{P(B)}$  and both numerator and denominator are never negative.

Axiom 2:  $P(S|B) = 1$ , because  $P(S|B) = \frac{P(S \cap B)}{P(B)} = \frac{P(B)}{P(B)} = 1$ .

Axiom 3:  $P(A_1 \cup A_2 | B) = P(A_1 | B) + P(A_2 | B)$  if  $A_1$  and  $A_2$  are disjoint because

$$P(A_1 \cup A_2 | B) = \frac{P((A_1 \cup A_2) \cap B)}{P(B)} = \frac{P((A_1 \cap B) \cup (A_2 \cap B))}{P(B)}$$

$$\begin{aligned}
&= \frac{P(A_1 \cap B) + P(A_2 \cap B)}{P(B)} \quad \text{since } A_1 \cap B \text{ and } A_2 \cap B \text{ are disjoint} \\
&= \frac{P(A_1 \cap B)}{P(B)} + \frac{P(A_2 \cap B)}{P(B)} = P(A_1|B) + P(A_2|B)
\end{aligned}$$

Axiom 4:  $P(A_1 \cup A_2 \cup \dots | B) = P(A_1|B) + P(A_2|B) + \dots$  if  $A_1, A_2, \dots$  is sequence of disjoint events, because

$$\begin{aligned}
P(A_1 \cup A_2 \cup \dots | B) &= \frac{P((A_1 \cup A_2 \cup \dots) \cap B)}{P(B)} = \frac{P((A_1 \cap B) \cup (A_2 \cap B) \cup \dots)}{P(B)} \\
&= \frac{P(A_1 \cap B) + P(A_2 \cap B) + \dots}{P(B)} \quad \text{since } A_1 \cap B, A_2 \cap B, \dots \text{ are disjoint} \\
&= \frac{P(A_1 \cap B)}{P(B)} + \frac{P(A_2 \cap B)}{P(B)} + \dots = P(A_1|B) + P(A_2|B) + \dots
\end{aligned}$$

9. 2-50, p. 77, L-G.

$$\begin{aligned}
P(A \cap B \cap C) &= P(A \cap D) \quad \text{where } D = B \cap C \\
&= P(A|D) P(D) \quad \text{by the definition of conditional probability} \\
&= P(A|B \cap C) P(B \cap C) \quad \text{by definition of } d \\
&= P(A|B \cap C) P(B|C) P(C) \quad \text{by the definition of conditional probability}
\end{aligned}$$

10. 2-51, p. 77, L-G.

Probability model for the experiment:

$(N_1, N_2)$ , where  $N_1$  and  $N_2$  are outcomes of first and second dice, respectively

$$S = \{ (1,1), (1,2), \dots, (1,6), (2,1), \dots, (6,6) \} = \{ (i,j) : i \in \{1,2,\dots,6\}, j \in \{1,2,\dots,6\} \}$$

$$P((i,j)) = \frac{1}{36}$$

$$A = \text{"}N_1 + N_2 \text{ is even"} = \{ (i,j) : i \in \{1,2,\dots,6\}, j \in \{1,2,\dots,6\}, i + j \in \{2,4,\dots,12\} \}$$

$$B = \text{"}N_1 \text{ is even and } N_2 \text{ is even"} = \{ (i,j) : i \text{ is even, } j \text{ is even} \}$$

notice that  $B \subset A$

$$(a) \mathbf{P(A|B)} = \frac{P(A \cap B)}{P(B)} = \frac{P(B)}{P(B)} = \mathbf{1}$$

$$(b) \mathbf{P(B|A)} = \frac{P(B \cap A)}{P(A)} = \frac{P(B)}{P(A)} = \frac{|B|/36}{|A|/36} = \frac{|B|}{|A|} = \frac{3 \times 3}{6 \times 3} = \mathbf{\frac{1}{2}}$$

11. 2-56, p. 78, L-G

Probability model:

T = time of arrival

S = (8,9) (the interval)

$$P(A) = \frac{\text{length of A in hours}}{1 \text{ hour}} = \text{length of A in hours}$$

Let B = event she has not arrived by 8:30 = event she arrives after 8:30 = (8.5,9) (the interval)

Let A = event she arrives in between 8:30 and 8:31 = (8.5, 8.5+1/60) .

$$\text{Then } P(A|B) = \frac{P(A \cap B)}{P(B)} = \frac{P((8.5, 8.5+1/60))}{P((8.5, 9))} = \frac{1/60}{1/2} = \frac{1}{30}$$

Let B = event she has not arrived by 8:50 = event she arrives after 8:50 = (8 5/6,9)

Let A = event she arrives in between 8:50 and 8:31 = (8 5/6, 8 5/6+1/60) .

$$\text{Then } P(A|B) = \frac{P(A \cap B)}{P(B)} = \frac{P((8 \frac{5}{6}, 8 \frac{5}{6}+1/60))}{P((8 \frac{5}{6}, 9))} = \frac{1/60}{1/6} = \frac{1}{10}$$

Knowledge that the professor arrives by 9 AM implies that the probability of imminent arrival approaches 1 as time approaches 9 AM.

12. A store opens at time 0 and never closes. The probability that the first customer to enter the store arrives between times  $t_1$  and  $t_2$  is

$$\int_{t_1}^{t_2} e^{-t} dt, \text{ where } 0 \leq t_1 \leq t_2 < \infty. \text{ (Time is measured continuously from 0.)}$$

(a) Find the probability that the first customer arrives between times  $t$  and  $t+1$  given that no customer arrive at or before time  $t$ .

Let A = event that first customer arrives between  $t$  and  $t+1 = (t, t+1)$

Let B = event that no customer arrives before time  $t$

= event that first customer arrives at time  $t$  or later =  $[t, \infty)$

$$\begin{aligned} P(A|B) &= \frac{P(A \cap B)}{P(B)} = \frac{P((t, t+1) \cap [t, \infty))}{P([t, \infty))} = \frac{P((t, t+1))}{P([t, \infty))} = \frac{\int_t^{t+1} e^{-u} du}{\int_t^{\infty} e^{-u} du} \\ &= \frac{e^{-t} - e^{-t-1}}{e^{-t}} = 1 - e^{-1} \end{aligned}$$

(b) Discuss the dependence of your answer on  $t$ .

Interestingly the answer does not depend on  $t$ . That is, the probability of arrival in the next minute does not get larger as time passes. This is much different than the situation in the previous problem.

13. 2-57a., p. 78, L-G

Let X denote the input and Y the output.

(a)  $P(Y=0) = P(Y=0|X=0) P(X=0) + P(Y=0|X=1) P(X=1)$  by the law of total probability

$$= (1 - \epsilon_1) \frac{1}{2} + \epsilon_2 \frac{1}{2} = \frac{1}{2} (1 - \epsilon_1 + \epsilon_2)$$