1. 2-32, p. 76, L-G

The number of distinct ordered triplets $=60 \times 60 \times 60=\mathbf{6 0}^{\mathbf{3}}=\mathbf{2 1 6 , 0 0 0}$
2. $2-35$, p. 76 , L-G

The number of distinct 7-tuples $=8 \times 10 \times 10 \times 10 \times 10 \times 10 \times 10=8 \times \mathbf{1 0}^{6}$

3 2-38, p. 76, L-G
There are 3! permutations (possible orderings of 3 books) of which only one corresponds to the correct order. Assuming all permutations are equiprobable,
$P($ correct order $)=\frac{1}{3!}=\frac{1}{6}$
4. $2-40$, p. 77, L-G
number of ways of picking 1 out of $5=\binom{5}{1}=\mathbf{5}$
number of ways of picking 2 out of $5=\binom{5}{2}=\mathbf{1 0}$
number of ways of picking none, some or all $5=\sum_{i=1}^{5}\binom{5}{i}=\mathbf{2}^{5}$ (by binomial thm)
5. 2-43, p. 77, L-G

There are $\binom{49}{6}$ ways of choosing six numbers. Assuming each is equiprobable, the probability of winning is

$$
\frac{1}{\binom{49}{6}}=7.15 \times 10^{-8}
$$

6. 2-54, p. 78, L-G (Note: Corollary 1 is in Section 2.2.)

The situation described is equivalent to the experiment of choosing twenty numbers in sequence from the set $\{1,2, \ldots, 365\}$ with replacement and recording the 20 numbers in order. A probability model for the experiment is:

Variable representing the outcome: $\underline{X}=\left(\mathrm{X}_{1}, \ldots, \mathrm{X}_{20}\right)$.
Sample space: $S=\left\{\underline{x}=\left(x_{1}, \ldots, x_{20}\right): x_{i} \in\{1, \ldots, 365\}\right\}$
Probability law: $P(\underline{x})=\frac{1}{365^{20}}$ for $\underline{x} \in S$
(therefore, for any set $\mathrm{A}, \mathrm{P}(\mathrm{A})=|\mathrm{A}| \times \frac{1}{365^{20}}$ )
We are interested in the probability of the event A that 2 or more students have the same birthday. As suggested by the hint, it is easier to findthe probability of the complement $\mathrm{A}^{\mathrm{c}}$, which is the event that no two students have the same birthday.

$$
A^{c}=\left\{\underline{x}=\left(x_{1}, \ldots, x_{20}\right): x_{i} \in\{1, \ldots, 365\}, x_{i} \neq x_{j}, i=1, \ldots, 20, j=1, \ldots, 20\right\}
$$

The number of elements in $\mathrm{A}^{\mathrm{c}}$ is $365 \times 364 \times \ldots \times 346=\frac{365!}{345!}=1.037 \times 10^{51}$
And so $\mathrm{P}\left(\mathrm{A}^{\mathrm{c}}\right)=|\mathrm{A}| \times \frac{1}{365^{20}}=\frac{365!}{345!} \times \frac{1}{365^{20}}=.589$
$\Rightarrow \mathbf{P}(\mathrm{A})=1-\mathbf{P}\left(\mathrm{A}^{\mathrm{c}}\right)=.411$
7. A circular pan of water with radius 10 inches is stirred and the position of a floating spec is recorded after the water has come to rest.
(a) Determine the conditional probability that the spec is more than 1 inch from the center given that it is less than the 3 inces from the center.
We begin with a probability model:
Variable representing outcome: $(\mathrm{X}, \mathrm{Y})$ where X and Y represent the horizontal and vertical positions, respectively, of the spec after it comes to rest.
Sample space: $\mathrm{S}=\left\{(\mathrm{x}, \mathrm{y}): \mathrm{x}^{2}+\mathrm{y}^{2} \leq 10\right\}$
Probability law: $\mathrm{P}(\mathrm{A})=\frac{\operatorname{area} \mathrm{A}}{\operatorname{area~} \mathrm{S}}=\frac{\operatorname{area~A}}{100 \pi}$
(here we have given a general formula for the probability of any set)
We need to find $P(A \mid B)$ where $A=\left\{(x, y): 1<x^{2}+y^{2} \leq 10\right\}$ and $B=\left\{(x, y): x^{2}+y^{2}<3\right\}$
From the definition of conditional probability

$$
\mathbf{P}(\mathbf{A} \mid \mathbf{B})=\frac{\mathrm{P}(\mathrm{~A} \cap \mathrm{~B})}{\mathrm{P}(\mathrm{~B})}=\frac{\operatorname{area}(\mathrm{A} \cap \mathrm{~B}) / 100 \pi}{\operatorname{area}(\mathrm{~B}) / 100 \pi}=\frac{\operatorname{area}(\mathrm{A} \cap \mathrm{~B})}{\operatorname{area}(\mathrm{B})}=\frac{\pi 3^{2}-\pi 1^{2}}{\pi 3^{2}}=\frac{\mathbf{8}}{\mathbf{3}}
$$

(b) Determine the conditional probability that the spec is both to the north and east of the landing point, given that it is less than 5 inches from the center.
We need to find $P(A \mid B)$ where $A=\left\{(x, y): x^{2}+y^{2} \leq 10, x>0, y>0\right\}$ and $B=\left\{(x, y): x^{2}+y^{2}<5\right\}$ From the definition of conditional probability
$\mathbf{P}(\mathbf{A} \mid \mathbf{B})=\frac{\mathrm{P}(\mathrm{A} \cap \mathrm{B})}{\mathrm{P}(\mathrm{B})}=\frac{\operatorname{area}(\mathrm{A} \cap \mathrm{B}) / 100 \pi}{\operatorname{area}(\mathrm{~B}) / 100 \pi}=\frac{\operatorname{area}(\mathrm{A} \cap \mathrm{B})}{\operatorname{area}(\mathrm{B})}=\frac{\pi 5^{2} / 4}{\pi 5^{2}}=\frac{\mathbf{1}}{\mathbf{4}}$
8. 2-49, p. 77, L-G.

Let $B$ be a set such that $P(B)>0$, so the axioms of probability are satisfied. Then,
Axiom 1: $\mathrm{P}(\mathrm{A} \mid \mathrm{B}) \geq 0$ for every set A , because $\mathrm{P}(\mathrm{A} \mid \mathrm{B})=\frac{\mathrm{P}(\mathrm{A} \cap \mathrm{B})}{\mathrm{P}(\mathrm{B})}$ and both numerator and denominator are never negative.
Axiom 2: $\mathrm{P}(\mathrm{S} \mid \mathrm{B})=1$, because $\mathrm{P}(\mathrm{S} \mid \mathrm{B})=\frac{\mathrm{P}(\mathrm{S} \cap \mathrm{B})}{\mathrm{P}(\mathrm{B})}=\frac{\mathrm{P}(\mathrm{B})}{\mathrm{P}(\mathrm{B})}=1$.
Axiom 3: $\mathrm{P}\left(\mathrm{A}_{1} \cup \mathrm{~A}_{2} \mid \mathrm{B}\right)=\mathrm{P}\left(\mathrm{A}_{1} \mid \mathrm{B}\right)+\mathrm{P}\left(\mathrm{A}_{2} \mid \mathrm{B}\right)$ if $\mathrm{A}_{1}$ and $\mathrm{A}_{2}$ are disjoint because

$$
\begin{aligned}
\mathrm{P}\left(\mathrm{~A}_{1} \cup \mathrm{~A}_{2} \mid \mathrm{B}\right) & =\frac{\mathrm{P}\left(\left(\mathrm{~A}_{1} \cup \mathrm{~A}_{2}\right) \cap \mathrm{B}\right)}{\mathrm{P}(\mathrm{~B})}=\frac{\mathrm{P}\left(\left(\mathrm{~A}_{1} \cap \mathrm{~B}\right) \cup\left(\mathrm{A}_{2} \cap \mathrm{~B}\right)\right)}{\mathrm{P}(\mathrm{~B})} \\
& =\frac{\mathrm{P}\left(\mathrm{~A}_{1} \cap \mathrm{~B}\right)+\mathrm{P}\left(\mathrm{~A}_{2} \cap \mathrm{~B}\right)}{\mathrm{P}(\mathrm{~B})} \text { since } \mathrm{A}_{1} \cap \mathrm{~B} \text { and } \mathrm{A}_{2} \cap \mathrm{~B} \text { are disjoint }
\end{aligned}
$$

$$
=\frac{\mathrm{P}\left(\mathrm{~A}_{1} \cap \mathrm{~B}\right)}{\mathrm{P}(\mathrm{~B})}+\frac{\mathrm{P}\left(\mathrm{~A}_{2} \cap \mathrm{~B}\right)}{\mathrm{P}(\mathrm{~B})}=\mathrm{P}\left(\mathrm{~A}_{1} \mid \mathrm{B}\right)+\mathrm{P}\left(\mathrm{~A}_{2} \mid \mathrm{B}\right)
$$

Axiom 4: $P\left(A_{1} \cup A_{2} \cup \ldots \mid B\right)=P\left(A_{1} \mid B\right)+P\left(A_{2} \mid B\right)+\ldots \quad$ if $A_{1}, A_{2}, \ldots$ is sequence of disjoint events, because

$$
\begin{aligned}
& P\left(A_{1} \cup A_{2} \cup \ldots \mid B\right) \quad=\frac{P\left(\left(A_{1} \cup A_{2} \cup \ldots\right) \cap B\right)}{P(B)}=\frac{P\left(\left(A_{1} \cap B\right) \cup\left(A_{2} \cap B\right) \cup \ldots\right)}{P(B)} \\
&= \frac{P\left(A_{1} \cap B\right)+P\left(A_{2} \cap B\right)+\ldots}{P(B)} \text { since } A_{1} \cap B, A_{2} \cap B, \ldots \text { are disjoint } \\
&= \frac{P\left(A_{1} \cap B\right)}{P(B)}+\frac{P\left(A_{2} \cap B\right)}{P(B)}+\ldots=P\left(A_{1} \mid B\right)+P\left(A_{2} \mid B\right)+\ldots
\end{aligned}
$$

9. 2-50, p. 77, L-G.
$\mathrm{P}(\mathrm{A} \cap \mathrm{B} \cap \mathrm{C})=\mathrm{P}(\mathrm{A} \cap \mathrm{D})$ where $\mathrm{D}=\mathrm{B} \cap \mathrm{C}$
$=\mathrm{P}(\mathrm{A} \mid \mathrm{D}) \mathrm{P}(\mathrm{D})$ by the definition of conditional probability
$=P(A \mid B \cap C) P(B \cap C)$ by definition of $d$
$=P(A \mid B \cap C) P(B \mid C) P(C) \quad$ by the definition of conditional probability
10. 2-51, p. 77, L-G.

Probability model for the experiment:
$\left(\mathrm{N}_{1}, \mathrm{~N}_{2}\right)$, where $\mathrm{N}_{1}$ and $\mathrm{N}_{2}$ are outcomes of first and second dice, respectively $S=\{(1,1),(1,2), \ldots,(1,6),(2,1), \ldots,(6,6)\}=\{(i, j): i \in\{1,2, \ldots, 6\}, j \in\{1,2, \ldots, 6\}\}$
$P((i, j))=\frac{1}{36}$
$\mathrm{A}=\mathrm{"N}_{1}+\mathrm{N}_{2}$ is even" $=\{(\mathrm{i}, \mathrm{j}): \mathrm{i} \in\{1,2, \ldots, 6\}, \mathrm{j} \in\{1,2, \ldots, 6\}, \mathrm{i}+\mathrm{j} \in\{2,4, \ldots, 12\}\}$
$B=" N_{1}$ is even and $N_{2}$ is even" $=\{(i, j): i$ is even, $j$ is even $\}$
notice that $\mathrm{B} \subset \mathrm{A}$
(a) $\mathbf{P}(\mathbf{A} \mid \mathbf{B})=\frac{\mathrm{P}(\mathrm{A} \cap \mathrm{B})}{\mathrm{P}(\mathrm{B})}=\frac{\mathrm{P}(\mathrm{B})}{\mathrm{P}(\mathrm{B})}$
(b) $\mathbf{P}(\mathbf{B} \mid \mathbf{A})=\frac{\mathrm{P}(\mathrm{B} \cap \mathrm{A})}{\mathrm{P}(\mathrm{A})}=\frac{\mathrm{P}(\mathrm{B})}{\mathrm{P}(\mathrm{A})}=\frac{|\mathrm{B}| / 36}{|\mathrm{~A}| / 36}=\frac{|\mathrm{B}|}{|\mathrm{A}|}=\frac{3 \times 3}{6 \times 3}=\frac{\mathbf{1}}{\mathbf{2}}$
11. 2-56, p. 78, L-G

Probability model:

$$
\begin{aligned}
& \mathrm{T}=\text { time of arrival } \\
& \mathrm{S}=(8,9) \text { (the interval) } \\
& \mathrm{P}(\mathrm{~A})=\frac{\text { length of } \mathrm{A} \text { in hours }}{1 \text { hour }}=\text { length of } \mathrm{A} \text { in hours }
\end{aligned}
$$

Let $B=$ event she has not arrived by $8: 30=$ event she arrives after $8: 30=(8.5,9)$ (the interval)
Let $\mathrm{A}=$ event she arrives in between $8: 30$ and $8: 31=(8.5,8.5+1 / 60)$.
Then $\mathbf{P}(\mathbf{A} \mid \mathbf{B})=\frac{\mathrm{P}(\mathrm{A} \cap \mathrm{B})}{\mathrm{P}(\mathrm{B})}=\frac{\mathrm{P}((8.5,8.5+1 / 60))}{\mathrm{P}((8.5,9))}=\frac{1 / 60}{1 / 2}=\frac{\mathbf{1}}{\mathbf{3 0}}$
Let $B=$ event she has not arrived by $8: 50=$ event she arrives after $8: 50=(85 / 6,9)$
Let $A=$ event she arrives in between $8: 50$ and $8: 31=(85 / 6,85 / 6+1 / 60)$.
Then $\mathbf{P}(\mathbf{A} \mid \mathbf{B})=\frac{\mathrm{P}(\mathrm{A} \cap \mathrm{B})}{\mathrm{P}(\mathrm{B})}=\frac{\mathrm{P}((85 / 6,85 / 6+1 / 60))}{\mathrm{P}((85 / 6,9))}=\frac{1 / 60}{1 / 6}=\frac{\mathbf{1}}{\mathbf{1 0}}$
Knowledge that the professor arrives by 9 AM implies that the probability of imminent arrival approaches 1 as time approaches 9 AM.
12. A store opens at time 0 and never closes. The probability that the first customer to enter the store arrives between times $t_{1}$ and $t_{2}$ is

$$
\int_{t_{1}}^{t_{2}} e^{-t} d t \text {, where } 0 \leq t_{1} \leq t_{2}<\infty \text {. (Time is measured continuously from } 0 . \text { ) }
$$

(a) Find the probability that the first customer arrives between times $t$ and $t+1$ given that no customer arrive at or before time $t$.

Let $\mathrm{A}=$ event that first customer arrives between t and $\mathrm{t}+1=(\mathrm{t}, \mathrm{t}+1)$
Let $B=$ event that no customer arrives before time $t$
$=$ event that first customer arrives at time t or later $=[\mathrm{t}, \infty)$
$\mathbf{P}(\mathbf{A} \mid \mathbf{B})=\frac{P(A \cap B)}{P(B)}=\frac{P((t, t+1) \cap[t, \infty))}{P([t, \infty))}=\frac{P((t, t+1))}{P([t, \infty))}=\frac{\int_{t}^{t+1} e^{-u} d u}{\int_{t}^{\infty} e^{-u} d u}$

$$
=\frac{e^{-t-e^{-t-1}}}{e^{-t}}=\mathbf{1} \cdot \mathbf{e}^{-1}
$$

(b) Discuss the dependence of your answer on $t$.

Interestingly the answer does not depend on $t$. That is, the probability of arrival in the next minute does not get larger as time passes. This is much different than the situation in the previous problem.
13. 2-57a., p. 78, L-G

Let X denote the input and Y the output.
(a) $\mathbf{P}(\mathbf{Y}=0)=\mathrm{P}(\mathrm{Y}=0 \mid \mathrm{X}=0) \mathrm{P}(\mathrm{X}=0)+\mathrm{P}(\mathrm{Y}=0 \mid \mathrm{X}=1) \mathrm{P}(\mathrm{X}=1)$ by the law of total probability

$$
=\quad\left(1-\varepsilon_{1}\right) \frac{1}{2}+\varepsilon_{2} \frac{1}{2}=\frac{1}{2}\left(1-\varepsilon_{1}+\varepsilon_{2}\right)
$$

