

1. 4.30, p. 259

The pmf of  $Y$  given  $X = -1$  is  $p_{Y|X}(y|-1) = \frac{p_{XY}(x,y)}{p_Y(-1)}$  as given in the following table

	<b>Y = -1</b>	<b>0</b>	<b>1</b>
<b>i</b>	$\frac{1}{2}$	<b>0</b>	$\frac{1}{2}$
<b>ii</b>	$\frac{1}{3}$	$\frac{1}{3}$	$\frac{1}{3}$
<b>iii</b>	<b>0</b>	<b>0</b>	<b>1</b>

2. 4.32, case (iii), p. 259

$$f_{XY}(x,y) = 2, \quad 0 \leq x \leq 1, \quad 0 \leq y \leq 1-x. \quad f_X(x) = \int_{-\infty}^{\infty} f_{XY}(x,y) dy = \int_0^{1-x} 2 dx = 2-2x$$

$$f_{Y|X}(y|x) = \frac{f_{XY}(x,y)}{f_X(x)} = \begin{cases} \frac{2}{2-2x} = \frac{1}{1-x}, & 0 \leq x \leq 1, \quad 0 \leq y \leq 1-x \\ 0, & \text{else} \end{cases}$$

3. 4.35, p. 260

The problem statement gives us the conditional pdf of  $T$  given that the clerk is  $i$ . This is a kind of nonnumerical conditioning. Specifically, it tells us that

$$f_T(t|i) = \alpha_i e^{-\alpha_i t}, \quad t \geq 0,$$

The conditional cdf is  $F_T(t|i) = 1 - e^{-\alpha_i t}$

(a) To find the pdf of  $T$  we first find the cdf and then differentiate. Using the law of total probability

$$F_T(t) = \Pr(T \leq t) = \sum_{i=1}^n p_i \Pr(T \leq t|i) = \sum_{i=1}^n p_i F_T(t|i)$$

Taking the derivative gives

$$\begin{aligned} f_T(t) &= \frac{d}{dt} F_T(t) = \frac{d}{dt} \sum_{i=1}^n p_i F_T(t|i) = \sum_{i=1}^n p_i \frac{d}{dt} F_T(t|i) = \sum_{i=1}^n p_i f_T(t|i) \\ &= \sum_{i=1}^n p_i \alpha_i e^{-\alpha_i t}, \quad t \geq 0 \end{aligned}$$

Notice that the above is a kind of law of total probability for densities.

$$(b) \quad E[T] = \int_{-\infty}^{\infty} t f_T(t) dt = \int_0^{\infty} t \sum_{i=1}^n p_i \alpha_i e^{-\alpha_i t} dt = \sum_{i=1}^n p_i \int_0^{\infty} t \alpha_i e^{-\alpha_i t} dt = \sum_{i=1}^n p_i \frac{1}{\alpha_i}$$

$$E[T^2] = \int_{-\infty}^{\infty} t^2 f_T(t) dt = \int_0^{\infty} t^2 \sum_{i=1}^n p_i \alpha_i e^{-\alpha_i t} dt = \sum_{i=1}^n p_i \int_0^{\infty} t^2 \alpha_i e^{-\alpha_i t} dt = \sum_{i=1}^n p_i \frac{2}{\alpha_i^2}$$

$$\text{var}(T) = E[T^2] - (E[T])^2 = \sum_{i=1}^n p_i \frac{2}{\alpha_i^2} - \left( \sum_{i=1}^n p_i \frac{1}{\alpha_i} \right)^2$$

4. A random variable  $X$  is exponentially distributed with expected value 10. Let  $Y$  be a random variable whose conditional pdf given  $X = x$  is

$$f_{Y|X}(y|x) = \begin{cases} \frac{1}{10} e^{-(x-y)/10}, & y \geq x \\ 0, & \text{else} \end{cases}$$

Find the conditional pdf of  $X$  given  $Y = y$ .

By Bayes rule for conditional densities:  $f_{X|Y}(x|y) = \frac{f_{Y|X}(y|x)f_X(x)}{f_Y(y)}$

Since  $X$  is exponentially distributed with expected value 10:  $f_X(x) = \frac{1}{10} e^{-x/10}$ ,  $x \geq 0$ .

To find  $f_Y(y)$  we use  $f_Y(y) = \int_{-\infty}^{\infty} f_X(x) f_{Y|X}(y|x) dx$

$$f_X(x) f_{Y|X}(y|x) = \frac{1}{10} e^{-x/10} \frac{1}{10} e^{-(x-y)/10} = \frac{1}{100} e^{-y/10}, \quad 0 \leq x \leq y$$

$$\text{So ... } f_Y(y) = \int_0^y \frac{1}{100} e^{-y/10} dx = \frac{1}{100} y e^{-y/10}$$

Substituting into Bayes rule gives:

$$f_{X|Y}(x|y) = \frac{\frac{1}{100} e^{-y/10}}{\frac{1}{100} y e^{-y/10}} = \frac{1}{y}, \quad 0 \leq x \leq y$$

5. Joe buys his gasoline at the local cut-rate gas station. The station sells two grades of gas, A and B, but the station owner won't tell which you get. Joe is concerned because his car mileage varies uniformly between 20 and 32 miles per gallon with brand A and between 16 and 26 miles per gallon with brand B. Suppose Joe knows the probabilities of the gas station having brands A and B are .3 and .7, respectively. Let  $X$  be a random variable representing Joe's gas mileage.

(a) Find the cumulative distribution function of  $X$ .

From the problem statement, we deduce  $\Pr(\text{brand A gas})=0.3$ ,  $\Pr(\text{brand B gas})=0.7$ ; and

$$f_X(x|A) = \begin{cases} \frac{1}{12}, & 20 \leq x \leq 32 \\ 0, & \text{else} \end{cases} \quad f_X(x|B) = \begin{cases} \frac{1}{10}, & 16 \leq x \leq 26 \\ 0, & \text{else} \end{cases}$$

The cdf is

$$\begin{aligned} F_X(x) &= \Pr(X \leq x) = \Pr(X \leq x | A) \Pr(A) + \Pr(X \leq x | B) \Pr(B) \\ &= \int_{-\infty}^x f_X(x|A) dx \times 0.3 + \int_{-\infty}^x f_X(x|B) dx \times 0.7 \\ &= \begin{cases} 0, & x < 16 \\ 0.7 \int_{-16}^x \frac{1}{10} dx, & 16 \leq x < 20 \\ 0.3 \int_{-20}^x \frac{1}{12} dx + 0.7 \int_{-16}^x \frac{1}{10} dx, & 20 \leq x < 26 \\ 0.7 + 0.3 \int_{-20}^x \frac{1}{12} dx, & 26 \leq x < 32 \\ 1, & x \geq 32 \end{cases} = \begin{cases} 0, & x < 16 \\ 0.7 \frac{x-16}{10}, & 16 \leq x < 20 \\ 0.3 \frac{x-20}{12} dx + 0.7 \frac{x-16}{10} dx, & 20 \leq x < 26 \\ 0.7 + 0.3 \frac{x-20}{12} dx, & 26 \leq x < 32 \\ 1, & x \geq 32 \end{cases} \end{aligned}$$

(b) Joe has found that his gas mileage is at least 24. Given this, determine the probability that he is using brand B.

$$\begin{aligned} \Pr(\mathbf{B}|\mathbf{X}\geq 24) &= \frac{\Pr(\mathbf{X}\geq 24|\mathbf{B}) \Pr(\mathbf{B})}{\Pr(\mathbf{X}\geq 24)} = \frac{\int_{24}^{\infty} f_{\mathbf{X}}(x|\mathbf{B}) dx \times 0.7}{1 - F_{\mathbf{X}}(24)} = \frac{\int_{24}^{26} \frac{1}{10} dx \times 0.7}{1 - (0.3 \times 4/12 + 0.7 \times 8/10)} \\ &= \frac{\frac{1}{5} \times 0.7}{\frac{17}{50}} = \frac{\mathbf{7}}{\mathbf{17}} \end{aligned}$$

6. A gambler brings  $X$  dollars to a casino where  $X$  is a random variable with density

$$f_X(x) = \begin{cases} \frac{x}{80,000}, & 0 \leq x \leq 400 \\ 0, & \text{otherwise} \end{cases}$$

After a night of gambling the gambler leaves the casino with  $Y$  dollars, where  $Y$  is uniformly distributed between 0 and  $X$ .

(a) Given the gambler leaves the casino with less than \$200 dollars, find the probability that he brought less than \$200.

We deduce from the problem statement, that

$$f_{Y|X}(y|x) = \begin{cases} \frac{1}{x}, & 0 \leq y \leq x \\ 0, & \text{else} \end{cases}$$

and from this, we deduce that

$$f_{XY}(x,y) = f_X(x) f_{Y|X}(y|x) = \begin{cases} \frac{1}{80,000}, & 0 \leq x \leq 400, 0 \leq y \leq x \\ 0, & \text{else} \end{cases}$$

Now, we must find

$$\Pr(\mathbf{X} \leq 200 | \mathbf{Y} \leq 200) = \frac{\Pr(\mathbf{X} \leq 200, \mathbf{Y} \leq 200)}{\Pr(\mathbf{Y} \leq 200)}$$

The numerator:

$$\begin{aligned} \Pr(\mathbf{X} \leq 200, \mathbf{Y} \leq 200) &= \int_{-\infty}^{200} \int_{-\infty}^{200} f_{XY}(x,y) dy dx = \int_0^{200} \int_0^x \frac{1}{80,000} dy dx = \int_0^{200} \frac{x}{80,000} dx \\ &= \frac{1}{80,000} \frac{1}{2} x^2 \Big|_0^{200} = \frac{1}{4} \end{aligned}$$

The denominator:

$$\begin{aligned} \Pr(\mathbf{Y} \leq 200) &= \int_{-\infty}^{200} f_Y(y) dy = \int_{-\infty}^{200} \int_{-\infty}^{\infty} f_{XY}(x,y) dx dy = \int_0^{200} \int_y^{400} \frac{1}{80,000} dx dy \\ &= \int_0^{200} \frac{400-y}{80,000} dy = 1 - \frac{1}{80,000} \frac{1}{2} y^2 \Big|_0^{200} = \frac{3}{4} \end{aligned}$$

Substituting for the numerator and denominator gives

$$\Pr(\mathbf{X} \leq 200 | \mathbf{Y} \leq 200) = \frac{\mathbf{1}}{\mathbf{3}}$$

(b) Find the probability that his loss was less than \$100.

His loss is  $X - Y$ , so we need to find

$$\begin{aligned} \Pr(\mathbf{X-Y < 100}) &= \int_{-\infty}^{\infty} \int_{x-100}^x f_{XY}(x,y) dy dx = \int_0^{400} \int_{\max(0,x-100)}^x \frac{1}{80,000} dy dx \\ &= \int_0^{100} \int_0^x \frac{1}{80,000} dy dx + \int_{100}^{400} \int_{x-100}^x \frac{1}{80,000} dy dx \\ &= \int_0^{100} \frac{x}{80,000} dy + \int_{100}^{400} \frac{1}{800} dx = \frac{1}{80,000} \frac{1}{2} x^2 \Big|_0^{100} + \frac{300}{800} = \frac{1}{16} + \frac{3}{8} = \frac{7}{16} \end{aligned}$$

(c) Find the probability that his loss was exactly \$75.

$$\Pr(\mathbf{X-Y=75}) = \int_{-\infty}^{\infty} \int_{x-75}^{x-75} f_{XY}(x,y) dy dx = \mathbf{0} \text{ because the area of integration is zero}$$

(d) Find the density of  $Y$ .

$$\mathbf{f_Y(y)} = \int_{-\infty}^{\infty} f_{XY}(x,y) dx = \begin{cases} \int_y^{400} \frac{1}{80,000} dx, & 0 \leq y \leq 400 \\ y, & \text{else} \end{cases} = \begin{cases} \frac{400-y}{80,000}, & 0 \leq y \leq 400 \\ \mathbf{0}, & \text{else} \end{cases}$$

7. Suppose  $X$  and  $Y$  are jointly continuous, independent random variables. Show that

$$P(Y \leq X) = \int_{-\infty}^{\infty} F_Y(y) f_X(y) dy$$

$$\Pr(\mathbf{Y \leq X}) = \int_{-\infty}^{\infty} \int_{-\infty}^x f_{XY}(x,y) dy dx = \int_{-\infty}^{\infty} \int_{-\infty}^x f_X(x) f_Y(y) dy dx \text{ since } X \text{ and } Y \text{ are independent}$$

$$= \int_{-\infty}^{\infty} f_X(x) \int_{-\infty}^x f_Y(y) dy dx = \int_{-\infty}^{\infty} f_X(x) F_Y(x) dx = \int_{-\infty}^{\infty} f_X(y) F_Y(y) dy$$