Homework Set 9 DRAFT EECS 401

 Joe buys his gasoline at the local cut-rate gas station. The station sells two grades of gas, A and B, but the station owner won't tell which you get. Joe is concerned because his car mileage varies uniformly between 20 and 32 miles per gallon with brand A and between 16 and 26 miles per gallon with brand B. Suppose Joe knows the probabilities of the gas station having brands A and B are .3 and .7, respectively. Let X be a random variable representing Joe's gas mileage. (This is same descriptions as Prob. 5 of previous homework.)

Find the expected value of X WITHOUT integrating x times the density of X.

We use a version of the law of total expectation, namely,

$$E[X] = E[X|A] P(A) + E[X|B] P(B)$$

From the problem statement, we deduce Pr(A) = 0.3, Pr(B) = 0.7; and

$$f_{X}(x|A) = \begin{cases} \frac{1}{12}, \ 20 \le x \le 32\\ 0, \ else \end{cases} \qquad f_{X}(x|B) = \begin{cases} \frac{1}{10}, \ 16 \le x \le 26.\\ 0, \ else \end{cases}$$
  
Hence,  $E[X|A] = \int_{20}^{32} \frac{1}{12} dx = 26$  and  $E[X|B] = \int_{-16}^{26} \frac{1}{10} dx = 21$ , and

$$E[X] = 26 \times .3 + 21 \times .7 = 22.5$$

2. A gambler brings X dollars to a casino where X is a random variable with density

 $f_X(x) = \begin{cases} \frac{x}{80,000}, & 0 \le x \le 400\\ 0, & \text{otherwise} \end{cases}$ 

After a night of gambling the gambler leaves the casino with Y dollars, where Y is uniformly distributed between 0 and X. (This is same descriptions as Prob. 6 of previous homework.)

(a) Find the expected value of Y WITHOUT integrating x times the density of Y.

We use a version of the law of total expectation, namely,  $E[Y] = \int_{\infty}^{\infty} E[Y|X=x] f_X(x) dx$ 

From the problem statement, we deduce  $f_{Y|X}(y|x) = \begin{cases} \frac{1}{x}, & 0 \le y \le x \\ 0, & else \end{cases}$ 

Therefore,  $E[Y|X=x] = \int_{-\infty}^{\infty} y f_{Y|X}(y|x) dy = \int_{0}^{x} y \frac{1}{x} dy = \frac{1}{x} \frac{y^2}{2} \Big|_{0}^{x} = \frac{x}{2}$ 

 $\Rightarrow$ 

 $\mathbf{E[Y]} = \int_{0}^{400} \frac{x}{2} \frac{x}{80,000} \, dx = \frac{1}{160,000} \frac{x^3}{3} \Big|_{0}^{400} = \frac{400}{3} = \mathbf{133.3}$ 

(a) Given the gambler leaves the casino with \$200 dollars, find the expected number of dollars that he brought.

We need to find  $E[X|Y=200] = \int_{-\infty}^{\infty} x f_{X|Y}(x|200) dx$ 

To find  $f_{X|Y}(x|y)$  we use the following version of Bayes rule:  $f_{X|Y}(x|200) = \frac{f_{Y|X}(200|x) f_X(x)}{f_Y(200)}$ 

where  $f_Y(200) = \int_{-\infty}^{\infty} f_X(x) f_{Y|X}(200|x) dx = \int_{200}^{400} \frac{x}{80,000} \frac{1}{x} dx = \frac{200}{80,000} = \frac{1}{400}$ (note that  $f_{Y|X}(y|x) = \frac{x}{80,000}$  only if  $0 \le y \le x$ )

So 
$$f_{X|Y}(x|200) = \frac{\frac{1}{x} \frac{x}{80,000}}{\frac{1}{400}} = \frac{1}{200}$$
 if  $200 \le x \le 400$   
And  $\mathbf{E}[X|Y=200] = \int_{200}^{400} x \frac{1}{200} dx = \frac{1}{200} \frac{x^2}{2} \Big|_{200}^{400} = 400 - 100 = 300$ 

3. 4.39, p. 260

$$f_{X,Y,Z}(x,y,z) \ = \ k(x+y+z), \quad 0 \le x \le 1, \ 0 \le y \le 1, \ 0 \le z \le 1$$

(a) We choose k so that  $f_{X,Y,Z}(x,y,z)$  integrates to one.

$$1 = \int_{0}^{1} \int_{0}^{1} \int_{0}^{1} f_{X,Y,Z}(x,y,z) \, dz \, dy \, dx = \int_{0}^{1} \int_{0}^{1} \int_{0}^{1} k \, (x+y+z) \, dz \, dy \, dx$$
$$= \int_{0}^{1} \int_{0}^{1} k \, (x+y+\frac{1}{2}) \, dy \, dx = \int_{0}^{1} k \, (x+\frac{1}{2}+\frac{1}{2}) \, dx = k \, (\frac{1}{2}+\frac{1}{2}+\frac{1}{2}) = \frac{3}{2} \, k \implies \mathbf{k} = \frac{2}{3}$$
(b)  $f_{X,Y}(x,y) = \int_{0}^{1} f_{X,Y,Z}(x,y,z) \, dz = \frac{2}{3} \, (x+y+\frac{1}{2}), \ 0 \le x \le 1, \ 0 \le y \le 1$ 
$$\mathbf{f}_{Z|X,Y}(\mathbf{z}|\mathbf{x},\mathbf{y}) = \frac{f_{X,Y,Z}(x,y,z)}{f_{X,Y}(x,y)} = \frac{\frac{2}{3}(x+y+\frac{1}{2})}{\frac{2}{3}(x+y+\frac{1}{2})} = \frac{(\mathbf{x}+\mathbf{y}+\mathbf{z})}{(\mathbf{x}+\mathbf{y}+\frac{1}{2})}$$

4. 4.41, p. 261 (*don't use the chain rule; you are, in effect, rederiving it for 3 random variables.*) We need to show  $f_{X,Y,Z}(x,y,z) = f_X(x) f_{Y|X}(y|x) f_{Z|X,Y}(z|x,y)$ . Let us start with the right hand side and substitute the formulas for each of the conditional densities there. We have

$$f_X(x) \ f_{Y|X}(y|x) \ f_{Z|X,Y}(z|x,y) \ = \ f_X(x) \ \frac{f_{XY}(x,y)}{f_X(x)} \ \frac{f_{XYZ}(x,y,z)}{f_{XY}(x,y)} \ = \ f_{XYZ}(x,y,z)$$

Alternate derivation: We use the fact that  $f_{U,V}(u,v) = f_U(u) f_{V|U}(v|u)$ 

With U = (X,Y), u = (x,y), V = Z and v = z, the above fact implies

$$\begin{split} f_{X,Y,Z}(x,y,z) \ &= \ f_{U,V}(u,v) \ = \ f_{U}(u) \ f_{V|U}(v|u) \ = \ f_{X,Y}(x,y) \ f_{Z|X,Y}(z|x,y) \\ &= \ f_{X}(x) \ f_{Y|X}(y|x) \ f_{Z|X,Y}(z|x,y) \end{split}$$

where the last step follows by the above stated fact, with U = X, u = x, V = Y, v = y

5. Let W = X + 3 Y - Z, where X, Y and Z be uncorrelated Gaussian random variables with means  $m_x = 1$ ,  $m_Y = 2$  and  $m_Z = 3$ , and variances  $\sigma_X^2 = 1$ ,  $\sigma_Y^2 = 2$ ,  $\sigma_Z^2 = 3$ . Find  $P(W \ge 1)$ . Hint: W is a linear combination of Gaussian random variables.

W is Gaussian, since it is the linear combination of Gaussian random variables. Therefore, to find

the density of W we need only find E[W] and  $\sigma_W^2$ .

$$E[W] = E[X+3Y-Z] = E[X] + 3 E[Y] - E[Z] = 1 + 3 \times 2 - 3 = 4$$

Since X, Y and Z are uncorrelated, so are X, 3Y and -Z, so the variance of W = X + 3Y - Z is the sum of the variances of X, 3Y and -Z:

$$\sigma_{W}^{2} = var(X) + var(3Y) + var(-Z) = \sigma_{X}^{2} + 9\sigma_{Y}^{2} + \sigma_{Z}^{2} = 1 + 9 \times 2 + 3 = 22$$

Now,

$$P(W \ge 1) = \int_{1}^{\infty} f_{W}(w) dw = \int_{1}^{\infty} \frac{1}{\sqrt{2\pi\sigma_{W}^{2}}} \exp\left\{-\frac{(w-E[W])^{2}}{2\sigma_{W}^{2}}\right\} dw$$
  
=  $\int_{\frac{1-E[W]}{\sigma_{W}}} \frac{1}{\sqrt{2\pi\sigma_{W}^{2}}} \exp\left\{-\frac{u^{2}}{2}\right\} du \sigma_{W}$ , letting  $u = \frac{w-E[W]}{\sigma_{W}}$   
=  $Q\left(\frac{1-E[W]}{\sigma_{W}}\right) = Q\left(\frac{1-4}{\sqrt{22}}\right) = Q(-.64) = 1 - Q(.64) \cong 1 - .26 \cong .74$ 

6. 4.51, p. 262

 $f_{X,Y}(x,y) = 2e^{-(x+y)}, 0 \le y \le x$ , and Z = X + Y. Note that X and Y are not independent. As shown in class and in Example 4.31, p. 221,

$$f_{\mathbf{Z}}(\mathbf{z}) = \int_{\infty} f_{XY}(x,z-x) \, dx, \quad \text{note:} \ f_{XY}(x,z-x) = 2e^{-(x+z-x)} , \ 0 \le z-x \le x$$
$$= 2e^{-z}, \ z/2 \le x \le z = \int_{z/2}^{z} 2e^{-z} \, dx = 2e^{-z} \frac{z}{2} = \mathbf{z} e^{-z} , \ \mathbf{z} \ge \mathbf{0}$$

7. 4.61, p. 263

Since X and Y are independent  $X^2$  and Y are independent (functions of independent random variables are independent), so

 $E[X^2 Y] = E[X^2] E[Y]$ 

 $\infty$ 

Since X is Gaussian with E[X] = 0 and  $\sigma_X^2 = 1$ ,  $E[X^2] = \sigma_X^2 + (E[X])^2 = 1 + 0 = 1$ Since Y is uniform on [-1,3], E[Y] = 1.

$$E[X^2 Y] = E[X^2] E[Y] = 1 \times 1 = 1$$

8. 4.67, p. 263

$$\begin{split} Y &= aX + b \quad \text{and} \quad \rho = \frac{\text{cov}(X,Y)}{\sigma_X \sigma_Y} \ , \\ \text{cov}(X,Y) &= E[XY] - E[X] E[Y] = E[X(aX+b] - E[X] E[aX+b] \\ &= a E[X^2] + b E[X] - a (E[X])^2 - bE[X] = a (E[X^2] - (E[X])^2) = a \sigma_X^2 \end{split}$$

$$\sigma_{Y}^{2} = var(aX+b) = a^{2} \sigma_{X}^{2} \implies \rho = \frac{cov(X,Y)}{\sigma_{X}\sigma_{Y}} = \frac{a \sigma_{X}^{2}}{\sigma_{X}a\sigma_{X}} = 1$$

9. Let X,Y be continuous random variables with joint density

$$f_{XY}(x,y) = \begin{cases} x+y, \ 0 \le x \le 1, 0 \le y \le 1\\ 0, \ else \end{cases}$$

(a) Find the best linear estimate for Y given X=x.

The best linear estimate for Y given X = x, is

$$\begin{split} \hat{\mathbf{y}} &= \mathbf{a}\mathbf{x} + \mathbf{b} = \rho_{XY} \frac{\sigma_Y}{\sigma_X} (\mathbf{X} - \mathbf{E}[\mathbf{X}]) + \mathbf{E}[\mathbf{Y}] \\ \mathbf{f}_{\mathbf{Y}}(\mathbf{y}) &= \int_{-\infty}^{\infty} \mathbf{f}_{XY}(\mathbf{x}, \mathbf{y}) \, d\mathbf{x} = \int_{0}^{1} (\mathbf{x} + \mathbf{y}) \, d\mathbf{x} = \left(\frac{\mathbf{x}^2}{2}\Big|_{0}^{1} + \mathbf{y}\right) = \frac{1}{2} + \mathbf{y}, \ 0 < \mathbf{y} < 1 \\ \mathbf{f}_{\mathbf{X}}(\mathbf{x}) &= \int_{-\infty}^{\infty} \mathbf{f}_{\mathbf{X}Y}(\mathbf{x}, \mathbf{y}) \, d\mathbf{y} = \int_{0}^{1} (\mathbf{x} + \mathbf{y}) \, d\mathbf{y} = \left(\frac{\mathbf{y}^2}{2}\Big|_{0}^{1} + \mathbf{x}\right) = \frac{1}{2} + \mathbf{x}, \ 0 < \mathbf{x} < 1 \\ \mathbf{E}[\mathbf{X}] &= \mathbf{E}[\mathbf{Y}] = \int_{0}^{1} \mathbf{y} \left(\frac{1}{2} + \mathbf{y}\right) \, d\mathbf{y} = \frac{1}{4} + \frac{1}{3} = \frac{7}{12} \\ \mathbf{E}[\mathbf{X}^2] &= \mathbf{E}[\mathbf{Y}^2] = \int_{0}^{1} \mathbf{y}^2 \left(\frac{1}{2} + \mathbf{y}\right) \, d\mathbf{y} = \frac{1}{6} + \frac{1}{4} = \frac{5}{12} \\ \mathbf{E}[\mathbf{X}\mathbf{Y}] &= \int_{0}^{1} \int_{0}^{1} \mathbf{x} \mathbf{y} \left(\mathbf{x} + \mathbf{y}\right) \, d\mathbf{x} \, d\mathbf{y} = \int_{0}^{1} \frac{\mathbf{y}}{3} + \frac{\mathbf{y}^2}{2} \, d\mathbf{y} = \frac{1}{6} + \frac{1}{6} = \frac{1}{3} \\ \rho_{\mathbf{X}\mathbf{Y}} &= \frac{\mathbf{E}[\mathbf{X}\mathbf{Y}] - \mathbf{E}[\mathbf{X}] \mathbf{E}[\mathbf{Y}]}{\sqrt{\mathbf{E}[\mathbf{X}^2] - (\mathbf{E}[\mathbf{X}])^2} \sqrt{\mathbf{E}[\mathbf{Y}^2] - (\mathbf{E}[\mathbf{Y}])^2}} = \frac{\frac{1/3 - 49/144}{5/12 - 49/144} = -\frac{1}{11} \\ \mathbf{S}_0, \quad \hat{\mathbf{y}} &= -\frac{1}{11} \left(\mathbf{x} - \frac{7}{12}\right) + \frac{7}{12} \end{split}$$

(b) Find the best overall estimate for Y given X = x.

The best overall estimate for Y given X=x is  $\hat{y} = E[Y|X=x] = \int_{\infty}^{\infty} y f_{Y|X}(y|x) dx$ Using Bayes rule:  $f_{Y|X}(y|x) = \frac{f_{XY}(x,y)}{f_X(x)} = \frac{x+y}{1/2+x} = \frac{2(x+y)}{1+2x}, \ 0 < x < 1, \ 0 < y < 1$ Therefore,  $\hat{y} = \int_{\infty}^{\infty} y f_{Y|X}(y|x) dx = \int_{0}^{1} y \frac{2(x+y)}{1+2x} dy = \frac{2x}{1+2x} \int_{0}^{1} y dy + \frac{2}{1+2x} \int_{0}^{1} y^2 dy$   $= \frac{2x}{1+2x} \frac{y^2}{2} \Big|_{0}^{1} + \frac{2}{1+2x} \frac{y^3}{3} \Big|_{0}^{1}$  $= \frac{2}{1+2x} (\frac{x}{2} + \frac{1}{3})$ 

The two estimation rule are plotted  $\Rightarrow$  (One is linear, one is not.)

