

**PROBLEM:** GIVEN  $\underline{x}, \underline{y}$  ARE 2 VECTOR JOINTLY GAUSSIAN RANDOM VARIABLES,

PROVE  $\hat{\underline{x}}_{MSE}(\underline{y}) = E[\underline{x} | \underline{y} = \underline{y}] = \boxed{E(\underline{x}) + K_{xy} K_y^{-1} (\underline{y} - E(\underline{y}))}$

WHERE  $K_{xy} = E[(\underline{x} - E(\underline{x}))(\underline{y} - E(\underline{y}))^T]$  AND  $K_y = K_{yy}$ .

**PROOF:** WE SHOW  $f_{\underline{x}|\underline{y}}(\underline{x}|\underline{y}) = \frac{1}{(2\pi)^{n/2} |K_{x|y}|^{1/2}} e^{-\frac{1}{2} (\underline{x} - E(\underline{x}|\underline{y}))^T K_{x|y}^{-1} (\underline{x} - E(\underline{x}|\underline{y}))}$

WHERE  $E(\underline{x}|\underline{y}) = E(\underline{x}) + K_{xy} K_y^{-1} (\underline{y} - E(\underline{y}))$  AND  $K_{x|y} = K_x - K_{xy} K_y^{-1} K_{xy}^T$ .

THE DESIRED RESULT FOLLOWS IMMEDIATELY. NOTE  $E[\underline{e}\underline{e}^T] = \underline{I}$  (TAKE  $E_y(\cdot)$ )

LET  $\underline{z} = \begin{bmatrix} \underline{x} \\ \underline{y} \end{bmatrix}$  SO THAT  $E(\underline{z}) = \begin{bmatrix} E(\underline{x}) \\ E(\underline{y}) \end{bmatrix}$  AND  $K_z = \begin{bmatrix} K_x & K_{xy} \\ K_{xy}^T & K_y \end{bmatrix}$ .

LET  $\underline{z}' = \begin{bmatrix} \underline{I} & -A \\ 0 & \underline{I} \end{bmatrix} \underline{z}$ , WHERE  $A$  IS TO BE SPECIFIED. LET  $\underline{z}' = \begin{bmatrix} \underline{x}' \\ \underline{y} \end{bmatrix}$  (NOTE  $\underline{y}' = \underline{y}$ ).

THEN  $K_{z'} = \begin{bmatrix} \underline{I} & -A \\ 0 & \underline{I} \end{bmatrix} \begin{bmatrix} K_x & K_{xy} \\ K_{xy}^T & K_y \end{bmatrix} \begin{bmatrix} \underline{I} & 0 \\ -A^T & \underline{I} \end{bmatrix} = \begin{bmatrix} K_x - AK_{xy}^T - K_{xy}A^T + AK_yA^T & K_{xy} - AK_y \\ K_{xy}^T - K_yA^T & K_y \end{bmatrix}$

∴ CHOOSE  $A = K_{xy} K_y^{-1}$  TO UNCORRELATE  $\underline{x}'$  AND  $\underline{y}$ .

THIS CHOICE OF  $A \rightarrow E(\underline{x}') = E(\underline{x}) - AE(\underline{y}) = E(\underline{x}) - K_{xy} K_y^{-1} E(\underline{y})$  AND

$K_{x'} = K_x - AK_{xy}^T - K_{xy}A^T + AK_yA^T = K_x - K_{xy} K_y^{-1} K_{xy}^T$ .

NOW,  $f_z(\underline{z}) = f_{\underline{x}, \underline{y}}(\underline{x}, \underline{y}) = f_{\underline{x}', \underline{y}}(\underline{x}', \underline{y}) = f_{z'}(\underline{z}')$  SINCE  $\text{DET} \begin{bmatrix} \underline{I} & -A \\ 0 & \underline{I} \end{bmatrix} = 1$ .

THUS  $f_{\underline{x}|\underline{y}}(\underline{x}|\underline{y}) = f_{\underline{x}, \underline{y}}(\underline{x}, \underline{y}) / f_{\underline{y}}(\underline{y}) = f_{\underline{x}', \underline{y}}(\underline{x}', \underline{y}) / f_{\underline{y}}(\underline{y}) = f_{\underline{x}'|\underline{y}}(\underline{x}'|\underline{y}) = f_{\underline{x}'}(\underline{x}')$ .

∴  $f_{\underline{x}|\underline{y}}(\underline{x}|\underline{y}) = f_{\underline{x}'}(\underline{x}')$

$= \frac{1}{(2\pi)^{n/2} |K_{x'}|^{1/2}} e^{-\frac{1}{2} (\underline{x}' - E(\underline{x}'))^T K_{x'}^{-1} (\underline{x}' - E(\underline{x}'))}$

$= \frac{1}{(2\pi)^{n/2} |K_{x'}|^{1/2}} e^{-\frac{1}{2} (\underline{x} - \underbrace{[E(\underline{x}) + K_{xy} K_y^{-1} (\underline{y} - E(\underline{y}))]}_{= E(\underline{x}|\underline{y})})^T K_{x'}^{-1} (\underline{x} - \underbrace{[E(\underline{x}) + K_{xy} K_y^{-1} (\underline{y} - E(\underline{y}))]}_{= E(\underline{x}|\underline{y})})}$

SINCE  $K_{x'} = K_{x|y}$  (DEFINED ABOVE), WE ARE DONE.

SINCE  $\underline{x}'$  AND  $\underline{y}$  ARE UNCORRELATED → INDEPENDENT SINCE JOINTLY GAUSSIAN.

NOTE THAT  $\hat{\underline{x}}(\underline{y}) = E(\underline{x}) + K_{xy} K_y^{-1} (\underline{y} - E(\underline{y}))$  IS BEST LINEAR LEAST-SQUARES ESTIMATE OF  $\underline{x}$  GIVEN  $\underline{y} = \underline{y}$ . THIS CAN BE DERIVED AS IN SCALAR CASE, ALTHOUGH MESSIER