

DEF: A *discrete* random variable (rv) is a mapping $x : \Omega \rightarrow$ *countable* set.

DEF: A *continuous* random variable is a mapping $x : \Omega \rightarrow$ *uncountable* set.

WLOG: Countable range of discrete rv is 1-1 with \mathcal{Z} =integers.

DEF: The *pmf* (probability mass function) for a *discrete* rv x is

Props: $p_x(X) = Pr[x = X]$. We have: $0 \leq p_x(X) \leq 1$; $\sum_{X=-\infty}^{\infty} p_x(X) = 1$.
For *continuous* rvs, use pdf (probability density function) $f_x(X)$.

EX1: Discrete rv x =#heads in 3 independent flips of a fair coin.

$$p_x(0) = p_x(3) = \frac{1}{8}; \quad p_x(1) = p_x(2) = \frac{3}{8}; \quad p_x(X) = 0 \text{ otherwise.}$$

EX2: *Geometric pmf:* $p_x(X) = (1 - a)^{X-1}a, X \geq 1, \quad 0 < a < 1$.

Note: $\sum_{X=1}^{\infty} (1 - a)^{X-1}a = a/(1 - (1 - a)) = 1$ and $0 < (1 - a)^{X-1}a < 1$.

Note: PDF (distribution) $F_x(X)$ for discrete rv x is piecewise constant (staircase) function with a countable number of discontinuities.

But: Flip a fair coin. If heads, flip a 2^{nd} coin with $Pr[\text{heads}]=2/3$.

If 2^{nd} coin heads, set rv $x=3$; if 2^{nd} coin tails, set rv $x=1$.

If 1^{st} coin tails, spin wheel of fortune and multiply the result by 5.

1. Sample space $\Omega = (\{\text{heads}\} \times \{1, 3\}) \cup (\{\text{tails}\} \times [0, 5))$.
2. PDF is discontinuous (below left); pdf contains impulses (below right).
3. x is called a *mixture* rv: neither discrete nor continuous.
4. To express using pdf, need impulse to put finite prob. at a single point.
5. Example: $Pr[x \leq 2] = F_x(2) = \frac{1}{2}(1/3) + \frac{1}{2}(1/5)2 = 11/30$.
6. Note $Pr[x < 1] = 0.1 \neq Pr[x \leq 1] = 0.2666\dots$
due to the discontinuity of the PDF $F_x(X)$ at $x = 1$.

DEF: Joint PDF $F_{x_1 \dots x_N}(X_1 \dots X_N) = Pr[x_1 \leq X_1 \& x_2 \leq X_2 \dots]$.

DEF: Joint pdf $f_{x_1 \dots x_N}(X_1 \dots X_N) = \frac{\partial}{\partial X_1} \dots \frac{\partial}{\partial X_N} F_{x_1 \dots x_N}(X_1 \dots X_N)$.

Note: $Pr[X_1 < x_1 \leq X_1 + \delta X_1 \text{ AND } \dots X_N < x_N \leq X_N + \delta X_N]$
 $= f_{x_1 \dots x_N}(X_1 \dots X_N) \delta X_1 \dots \delta X_N$ (compare to 1-D pdf).

EX: Amy and Bo meet for lunch. 1st to arrive will wait only 15 minutes.
 Amy and Bo equally likely to arrive anytime between noon and 1 PM.

Q1: Assume arrival times independent. $Pr[\text{Amy and Bo meet for lunch}] = ?$

Let: $x = \text{Amy's arrival time in fraction of hour after noon}$; $y = \text{Bo's arrival}$.

Ω : Sample space: $\Omega = [0, 1]^2$. $f_{x,y}(X, Y) = 1$ for $(X, Y) \in [0, 1]^2$.

Event: $A = \text{they meet for lunch} = \{(X, Y) : |X - Y| < \frac{1}{4}\}$. Want $Pr[A]$.

$$Pr[A] = \underbrace{\int \int}_{|X-Y| < 1/4} f_{x,y}(X, Y) dX dY = \text{Area}[|X - Y| < 1/4] = 7/16.$$

Q2: Now suppose Amy and Bo are both more likely to arrive late:

$$f_x(X) = \begin{cases} 2X & \text{if } 0 < X < 1 \\ 0 & \text{otherwise} \end{cases}; \quad f_y(Y) = \begin{cases} 2Y & \text{if } 0 < Y < 1 \\ 0 & \text{otherwise} \end{cases}.$$

$$Pr[A] = \underbrace{\int \int}_{|X-Y| < 1/4} f_{x,y}(X, Y) dX dY = \underbrace{\int \int}_{|X-Y| < 1/4} (2X)(2Y) dX dY$$

$$= 1 - 2 \int_{\frac{1}{4}}^1 dX \int_0^{X-\frac{1}{4}} dY 4XY = 1 - 2 \int_0^{\frac{3}{4}} dX \int_{X+\frac{1}{4}}^1 dY 4XY = 139/256.$$

DEF: The *marginal* pdfs for x and y are computed from joint pdf $f_{x,y}(X, Y)$:
 $f_x(X) = \int f_{x,y}(X, Y) dY$; $f_y(Y) = \int f_{x,y}(X, Y) dX$.

DEF: x and y are *independent* rvs if $f_{x,y}(X, Y) = f_x(X) f_y(Y)$.

1. If given that x, y independent (as above), can compute $f_{x,y}(X, Y)$.
2. Given $f_{x,y}(X, Y)$, can compute marginal pdfs $f_x(X), f_y(Y)$ and check whether $f_{x,y}(X, Y) = f_x(X) f_y(Y)$. If so, x, y are independent.
3. If support (nonzero region) of $f_{x,y}(X, Y)$ is not rectangular, we know immediately that x, y are *not* independent rvs.

DEF: $p_{x,y|A}(X, Y|A) = Pr \left[\begin{matrix} x=X \\ y=Y \end{matrix} | A \right] = \begin{cases} p_{x,y}(X, Y)/Pr[A] & \text{if } (X, Y) \in A \\ 0 & \text{if } (X, Y) \notin A \end{cases}$

EX: Let $p_{x,y}(X, Y) = \frac{1}{4}$ if $(X, Y) \in [1, 2]^2$. Let $A = \{(X, Y) : X + Y \geq 3\}$.
 $p_{x,y|A}(X, Y|A) = \frac{1}{3}$ if $(X, Y) \in \{(1, 2), (2, 1), (2, 2)\}$; 0 otherwise.

DEF: *Conditional marginal pmf* $p_{x|A}(X|A) = \sum_Y p_{x,y|A}(X, Y|A)$.

EX: Here, $p_{x|A}(1|A) = \frac{1}{3}$; $p_{x|A}(2|A) = \frac{2}{3}$; $p_{x|A}(X|A) = 0$ otherwise.

DEF: *Conditional pmf* $p_{x|y}(X|Y) = Pr[x = X|y = Y] = \frac{p_{x,y}(X, Y)}{\sum_X p_{x,y}(X, Y)}$.

EX: $p_{x|y}(1|2) = p_{x|y}(2|2) = \frac{1}{2}$; $p_{x|y}(X|Y) = 0$ otherwise.

DEF: x, y *conditionally independent* if $p_{x,y|A}(X, Y|A) = p_{x|A}(X|A)p_{y|A}(Y|A)$.

DEF: *Conditional pdf* $f_{x|y} = f_{x,y}(X, Y)/f_y(Y)$.

Note: $Pr[X < x \leq X + \delta X | Y < y \leq Y + \delta Y] = \frac{Pr[X < x \leq X + \delta X \ \& \ Y < y \leq Y + \delta Y]}{Pr[Y < y \leq Y + \delta Y]}$
 $\rightarrow f_{x|y}(X|Y)\delta X = f_{x,y}(X, Y)\delta X \delta Y / (f_y(Y)\delta Y) = f_{x,y}(X, Y)/f_y(Y)\delta X$

EX: $f_{x,y}(X, Y) = c$ if $0 < X < 1 \ \& \ X < Y < 4X$; otherwise $f_{x,y}(X, Y) = 0$.

Let: $A = \{(X, Y) : X + Y \leq 2\}$. Compute conditional marginal $f_{x|A}(X|A)$.

c: $1 = \int \int f_{x,y}(X, Y)dX dY = c \frac{1}{2}(4 \cdot 1 - 1 \cdot 1) = c \frac{3}{2} \rightarrow c = \frac{2}{3}$.

Pr[A]: $Pr[A] = c \frac{1}{2}(2 \cdot 2 - 2 \cdot 1 - 2 \cdot 0.4) = 0.4$. See figure below.

1. $f_{x,y|A}(X, Y|A) = \frac{2}{3}/0.4$ for $(X, Y) \in A$; 0 otherwise.
 Always compute $f_{x,y|A}(X, Y|A)$ *first*, then $f_{x|A}(X|A)$.
2. $f_{x|A}(X|A) = \int f_{x,y|A}(X, Y|A)dY = \int (\frac{2}{3}/0.4)dY$ for $(X, Y) \in A$.
3. $f_{x|A}(X|A) = \begin{cases} \frac{1}{0.4} \frac{2}{3}(4X - X) = 5X & \text{if } 0 < X < 0.4 \\ \frac{1}{0.4} \frac{2}{3}(2 - X - X) = \frac{10}{3}(1 - X) & \text{if } 0.4 < X < 1 \end{cases}$