

1a. Note $E[(x(n) - y(n))^2] = E[x(n)^2] + E[y(n)^2] - 2E[x(n)y(n)] = R_x(n, n) + R_y(n, n) - 2R_{xy}(n, n) = 0 \rightarrow x(n) = y(n)$ **with probability one** from #6 of Problem Set #5.

1b. Let $y(n) = x(n + T) - x(n)$. Then $E[y(n)] = E[x(n + T)] - E[x(n)] = 0 - 0 = 0$ and $\sigma_{y(n)}^2 = E[(x(n + T) - x(n))^2] = R_x(0) + R_x(0) - 2R_x(T) = 0 \rightarrow x(n + T) = x(n) \rightarrow x(n)$ periodic with probability one, using the result of #1a above.
Also, $R_x(n + T) = E[x(i)x(i + n + T)] = E[x(i)x(i + n)] = R_x(n) \rightarrow R_x(n)$ periodic.

NOTE: Chebyshev $\rightarrow Pr[y > \epsilon] \leq \frac{\sigma_y^2}{\epsilon^2} = 0 \rightarrow Pr[y = 0] = 1$ (\Leftrightarrow #6 of Problem Set #5).

2. $x(n) = \rho x(n-1) + w(n), x(0) = 0 \Leftrightarrow x(n) = \sum_{i=0}^{n-1} h(i)w(n-i)$ where $h(n) = \rho^n, n \geq 0$
Note that the upper limit is $(n-1)$ since $x(0) = 0$ and $w(n)$ is only defined for $n \geq 1$.

2a. $E[x(n)] = \sum_{i=0}^{n-1} \rho^i E[w(n-i)] = 0$. Plugging directly into lecture formulae:

2b. $K_x(m, n) = \sum_{i=0}^{n-1} \sum_{j=0}^{n-1} \rho^{i+j} K_w(m-n-i+j) = \sum_{i=0}^{n-1} \sum_{j=0}^{n-1} \rho^{i+j} \sigma_w^2 \delta(m-n-i+j)$

$$K_x(m, n) = \sigma_w^2 \sum_{j=0}^{n-1} \rho^{m-n+2j} = \sigma_w^2 \rho^{m-n} \frac{1-\rho^{2n}}{1-\rho^2} = \sigma_w^2 \frac{\rho^{m-n} - \rho^{m+n}}{1-\rho^2} \text{ assuming } m \geq n.$$

$$K_x(m, n) = \sigma_w^2 \frac{\rho^{|m-n|} - \rho^{m+n}}{1-\rho^2} \text{ since } K_x(m, n) \text{ symmetric (or just redo for } m \leq n).$$

2c. For $|\rho| < 1$, clearly $\lim_{m, n \rightarrow \infty} K_x(m, n) = \frac{\sigma_w^2 \rho^{|m-n|}}{1-\rho^2}$. Hence $x(n)$ is *asymptotically* WSS.

This makes sense: the initial condition $x(0) = 0$ is now in the infinite past.

2d. As $\rho \rightarrow 1 \Leftrightarrow e = 1 - \rho \rightarrow 0$, then $K_x(m, n) = \sigma_w^2 \frac{\rho^{|m-n|} - \rho^{m+n}}{1-\rho^2} \rightarrow \sigma_w^2 \frac{(1-e)^{|m-n|} - (1-e)^{m+n}}{1-(1-e)^2}$
 $\rightarrow \sigma_w^2 \frac{e((m+n) - |m-n|)}{2e} = \sigma_w^2 \min[m, n]$, which is the II process result from lecture!

3. **N(n):** $N(n)$ is the sum of a 1-sided iid random process (Bernoulli), so have $N(n) = \sum_{k=1}^n x(k) \rightarrow E[N(n)] = nE[x] = np$ and $K_N(i, j) = \sigma_x^2 \min[i, j] = p(1-p) \min[i, j]$.

Y(n): $E[Y(n)] = E[(-1)^{N(n)}] = E[\prod_{k=1}^n (-1)^{x(k)}] = \prod_{k=1}^n E[(-1)^{x(k)}] = (1-2p)^n$
since $x(n)$ independent $\rightarrow (-1)^{x(n)}$ uncorrelated and $(-1)^{x(n)} = \begin{cases} 1 & \text{with prob. } 1-p \\ -1 & \text{with prob. } p \end{cases}$.

$E[Y(n)Y(n+k)] = E[\prod_{i=1}^n (-1)^{x(i)} \prod_{j=1}^{n+k} (-1)^{x(j)}] = E[\prod_{j=n+1}^{n+k} (-1)^{x(j)}] = (1-2p)^k$,
for $k > 0$, since $(-1)^{2x(j)} = 1$. Hence $R_Y(k) = (1-2p)^{|k|}$ (use same idea for $k < 0$).

But $K_Y(i, j) = R_Y(|i-j|) - E[Y(i)]E[Y(j)] = (1-2p)^{|i-j|} - (1-2p)^{(i+j)}$.

$Y(n)$ is *asymptotically* WSS since $\lim_{n \rightarrow \infty} (1-2p)^n = 0$. Note: Given $p \neq \frac{1}{2}$.

4. **Pmf for x(n):** $\{p_{x(n)}(n) = \frac{1}{n}; \quad p_{x(n)}(0) = 1 - \frac{1}{n}\} \rightarrow E[x(n)^2] = \frac{1}{n}n^2 + (1 - \frac{1}{n})0^2 = n$.

4a. Convergence in prob. $\Leftrightarrow \lim_{n \rightarrow \infty} Pr[|x(n) - 0| > \epsilon] = \lim_{n \rightarrow \infty} Pr[x(n) = 1] = \lim_{n \rightarrow \infty} \frac{1}{n} = 0$.

4b. Convergence with prob. $1 \Leftrightarrow Pr[\{s : \lim_{n \rightarrow \infty} x(n, s) = 0\}] = 1$. Fix $s \neq 0$ where $s \in \Omega$.
Then $x(n, s) = 0$ for $n > \frac{1}{s} \rightarrow \lim_{n \rightarrow \infty} x(n, s) = 0$. $Pr[\{s : s \neq 0\}] = 1 \rightarrow$ converges a.s.

4c. Convergence in mean square $\Leftrightarrow \lim_{n \rightarrow \infty} E[(x(n) - 0)^2] = \lim_{n \rightarrow \infty} n \neq 0 \rightarrow$ doesn't converge.

5. Note: $E[M(n)x(n)] = \frac{1}{n} \sum_{i=1}^n E[x(i)x(n)] = \frac{1}{n} \sum_{i=1}^n R_x(i, n) = C(n)$ since 0-mean.

Only if: Schwarz \neq : $C(n)^2 = E[M(n)x(n)]^2 \leq E[M(n)^2]E[x(n)^2] = E[M(n)^2]R_x(n, n)$.

Then $\stackrel{\text{L.I.M.}}{n \rightarrow \infty} M(n) = 0 \Leftrightarrow \stackrel{\text{LIM}}{n \rightarrow \infty} E[(M(n) - 0)^2] = 0 \rightarrow \stackrel{\text{LIM}}{n \rightarrow \infty} C(n) \rightarrow 0$. QED.

If: $E[M(n)^2] = \frac{1}{n^2} \sum_{i=1}^n \sum_{j=1}^n E[x(i)x(j)] = \frac{1}{n^2} \sum_{i=1}^n \sum_{j=1}^n R_x(i, j)$
 $= \frac{2}{n^2} \sum_{i=1}^n \sum_{j=1}^i R_x(j, i) - \frac{1}{n^2} \sum_{i=1}^n \sigma_{x(i)}^2 = \frac{2}{n^2} \sum_{i=1}^n iC(i) - \frac{1}{n^2} \sum_{i=1}^n \sigma_{x(i)}^2$.

Then $\stackrel{\text{LIM}}{n \rightarrow \infty} C(n) = 0 \rightarrow \stackrel{\text{LIM}}{n \rightarrow \infty} E[M(n)^2] = 0 \rightarrow \stackrel{\text{L.I.M.}}{n \rightarrow \infty} M(n) = 0$ using the hint

and $\frac{1}{n^2} \sum_{i=1}^n \sigma_{x(i)}^2 \leq \frac{1}{n} \text{MAX}_i R_x(i, i) \rightarrow 0$ since finite $\sigma_{x(i)}^2$. QED.
